

CRRA UTILITY MAXIMIZATION UNDER DYNAMIC RISK CONSTRAINTS

SANTIAGO MORENO-BROMBERG, TRAIAN A. PIRVU, AND ANTHONY RÉVEILLAC*

ABSTRACT. The problem of optimal investment with CRRA (constant, relative risk aversion) preferences, subject to dynamic risk constraints on trading strategies, is the main focus of this paper. Several works in the literature, which deal either with optimal trading under static risk constraints or with VaR-based dynamic risk constraints, are extended. The market model considered is continuous in time and incomplete, and the prices of financial assets are modeled by Itô processes. The dynamic risk constraints, which are time and state dependent, are generated by a general class of risk measures. Optimal trading strategies are characterized by a quadratic BSDE. Within the class of *time consistent distortion risk measures*, a three-fund separation result is established. Numerical results emphasize the effects of imposing risk constraints on trading.

1. Introduction

In this paper we consider the problem of a utility-maximizing agent, whose preferences are of constant relative risk aversion (CRRA) type, and whose trading strategies are subject to risk constraints. We work on a continuous-time, stochastic model with randomness being driven by Brownian noise. The market is incomplete and consists of several traded assets whose prices follow Itô processes.

In practice, managers set risk limits on the strategies executed by their traders. In fact, the mechanisms used to control risk are more complex: financial institution have specialized internal departments in charge of risk assessments. On top of that there are external regulatory institutions to whom financial institutions must periodically report their risk exposure. It is natural, therefore, to study the portfolio problem with risk constraints, which has received a great deal of scrutiny lately. A well known paper in this direction is [7]. The authors employ convex duality to characterize the optimal constrained portfolio. A more recent paper in the same direction is [13]. Here

Received 2012-2-29; Communicated by N. Privault.

2000 *Mathematics Subject Classification.* Primary 91B30; Secondary 60H30; 60G44.

Key words and phrases. BSDE, CRRA preferences, constrained utility maximization, correspondences, risk measures.

* Santiago Moreno-Bromberg acknowledges financial support from the Deutsche Forschungsgemeinschaft through the SFB 649 “Economic Risk” and from the Alexander von Humboldt Foundation via a research fellowship. The research leading to these results has received funding from the ERC (grant agreement 249415-RMAC) and from NCCR FinRisk (project Banking and Regulation). Traian A. Pirvu is grateful to NSERC through grant 5-36700 and MITACS through grant 5-26761, as well as SSHRC grant 5-26758.

Part of this work has been done while Anthony Réveillac was financed by the Deutsche Forschungsgemeinschaft Research center MATHEON.

the optimal constrained portfolio is characterized by a quadratic-growth Backward Stochastic Differential Equation (BSDE), which renders the method more amenable to numerical treatment. In these two (by now classical) papers the risk constraints are imposed either via abstract convex sets in [7] or via closed, time-independent sets in [13]. Lately, a line of research has been developed where the risk-constraint sets are specified employing a specific risk measure, e.g. VaR (Value-at-Risk) or TVaR (Tail Value-at-Risk). [6] introduced dynamically-consistent risk measurements in which the VaR and TVaR of a portfolio is reevaluated dynamically using the current information. In order to compute a portfolio dynamic VaR, TVaR, the authors assume that the distribution of the portfolio composition is kept unchanged during the period over which risk is assessed. In addition, the price is assumed to follow a linear coefficient with a deterministic volatility matrix and a deterministic drift coefficient. In such Markovian context the authors derive a characterization of the value function in terms of the solution to a Hamilton-Jacobi-Bellman PDE.

In this paper, we first extend the risk measurements introduced by [6] by considering a rich class of risk measures: we only require the latter to be Carathéodory maps. This class includes many convex and coherent risk measures. Then, we reduce the stochastic control problem with dynamic constraints to the study of a quadratic BSDE by employing the method developed in [13], which combines the martingale optimality principle with BSDE theory. The optimal, constrained trading strategies, as well as the value functions, can then be deduced from the solution of the BSDE. The main difference is that, unlike [13], our constraint sets are time dependent. This renders the methodology developed in [13] not directly applicable within our context. In addition, we derive a three-fund separation result for *time-consistent, distortion risk measures*. Finally, taking advantage of the BSDE formulation, we present some numerical results which allow one to evaluate the impact of constraints on the trading performance. We provide below a brief overview of the related literature.

Existing Research: A risk measure that is commonly used by both practitioners and academics is VaR. Despite its success, VaR has as drawbacks not being subadditive and not recognizing the accumulation of risk. This encouraged researchers to develop other risk measures, e.g. TVaR (Tail Value at Risk). The works on optimal investment with risk constraints generated by VaR, TVaR (or other risk measures) split into two categories, which depend on whether or not the risk assessment is performed in a static or a dynamic fashion. Let us briefly touch on the first category. The seminal paper is [3], where the optimal dynamic portfolio and wealth-consumption policies of utility maximizing investors who use VaR to control their risk exposure is analyzed. In a complete-market, Itô-processes framework, VaR is computed in a static manner (the authors compute the VaR of the final wealth only). An interesting finding is that VaR limits, when applied only at maturity, may actually increase risk. One way to overcome this problem is to consider a risk measure that is based on the risk-neutral expectation of loss - the Limited Expected Loss (LEL). In [10], a model with Capital-at-Risk (a version of VaR) limits in the Black-Scholes-Samuelson framework is presented. The authors assume that portfolio proportions are held constant during the whole investment period, which makes the problem static. [8] extends [10] from constant to deterministic parameters. In a market model with constant parameters,

[11] extends [3] to cover the case of bounded expected loss. In a general, continuous-time financial market model, [12] considers the portfolio problem under a downside risk constraint measured by an abstract convex risk measure. [16] extends [10] by imposing a uniform (in time) risk constraint.

In the category of dynamic risk measurements we recall the seminal paper [6]. Following the financial industry practice, the VaR (or some other risk measure) is computed (and dynamically re-evaluated) using a time window (2 weeks in practice) over which the trading strategies are assumed to be held constant for the purpose of risk measurement. The finding of the authors is that dynamic VaR and TVaR constraints reduce the investment (proportion wise) in the risky asset. [18] studies the impact of VaR constraint on equilibrium prices and the relationship with the leverage effect. [5] shows that, in equilibrium, VaR reduces market volatility. [23] finds that risk constraints may give rise to equilibrium asset-pricing bubbles. Among others, [2], [21], and [26] analyze the problem of investment and consumption subject to dynamic VaR constraints. [22] considers maximizing the growth rate of the portfolio in the context of dynamic VaR, TVaR and LEL constraints. In a complete market model, [24] uses a martingale method to study the optimal investment under dynamic risk constraints and partial information.

Our Contribution: This paper extends the risk measurements introduced by [6] by considering a relatively general class of risk measures (we only require them to be Carathéodory maps, a class that is rich enough to include many convex and coherent risk measures). The risk-constraint sets arising from such risk measures, and applied to the trading strategies, are time and state dependent. Moreover, they satisfy some important measurability properties. In addition, we go beyond the Markovian context of [6] by allowing the volatility and the drift coefficient in the dynamics of the price process to be non-deterministic and time-dependent.

We employ the methodology developed in [13] in order to provide existence of an optimal trading strategy subject to a dynamic risk constraint (see Theorem 3.5 which constitutes our first main result). Furthermore, we obtain a complete characterization of say optimal strategy and of the corresponding value function in terms of the solution of a quadratic growth BSDE (*c.f.* Theorem 3.6), which can be viewed as a generalization of the Hamilton–Jacobi–Bellman PDE in the Markovian realm. Our results depart from those in [13] in the sense that the constraint sets we consider are time dependent. This renders the methodology developed in [13] not directly applicable within our context. The difficulty stems from establishing the measurability of the BSDE’s driver (the BSDE that characterizes the optimal trading strategy). This is done by means of the Measurable Maximum Theorem and the Kuratowski–Ryll–Nardzewski Selection Theorem. After this step is achieved we apply results from [19] to get existence of solutions to the BSDE, which in turn yields the optimal trading strategy.

We then restrict our analysis to the class of *time consistent distortion risk measures*. By doing so we observe that the risk constraints have a particular structure: they are compact sets (for a fixed time and state) and they depend on two statistics (portfolio return and variance). This leads to a three-fund separation result. More precisely, an investor subject to regulatory constraints will invest her wealth into three-funds: a savings account and two index funds. One index fund is a mix of the stocks with

weights given by the Merton proportion. This index fund is related to market risk and most of the portfolio separation results refer to it. The second index is related to volatility risk. In a market with non-random drift and volatility the second index is absent. Thus, the second index can be explained by the demand of hedging volatility risk.

Numerical results shed light into the structure of the optimal trading strategy. More precisely, using recent results concerning numerical methods for quadratic growth BSDEs, we present in Section 5 some numerical examples for Value-at-Risk, Tail-Value-at-Risk and Limited-Expected-Loss. Our simulations clearly exhibit the effect of the risk constraint on the optimal strategy and on the associated value function: from the plots we observe that the risk constraints reduce gambling on the risky assets.

The paper is organized as follows: In Section 2 we introduce the basic model, the risk measures and the corresponding risk constraints. Section 3 presents measurability properties of the candidate optimal trading strategy and its characterization via a quadratic BSDE. In Section 4, *time consistent distortion risk measures* are considered. A three-fund separation result is obtained within this context. Numerical results are presented in Section 5. The paper ends with an appendix that contains some technical results.

2. Model Description and Problem Formulation

2.1. The financial market. Our model of a financial market, based on a filtered probability space $(\Omega, \mathcal{F}, \{\mathcal{F}_t\}_{t \in [0, T]}, \mathbb{P})$ that satisfies the usual conditions, consists of $n+1$ assets. The first one, $\{S_t^0\}_{t \in [0, T]}$, is a *riskless bond* with a strictly positive, constant interest rate $r > 0$. The remaining n assets are *stocks*, and their prices are modeled by a n -dimensional Itô-process $\{\mathbf{S}_t\}_{t \in [0, T]} = \{(S_t^i)_{i=1, \dots, n}\}_{t \in [0, T]}$. Their dynamics are given by the following stochastic differential equations, in which $\{\mathbf{W}_t\}_{t \in [0, T]} = \{(W_t^i)_{i=1, \dots, m}\}_{t \in [0, T]}$ is a m -dimensional standard Brownian motion:

$$\left. \begin{aligned} dS_t^0 &= S_t^0 r dt \\ dS_t^i &= S_t^i \left(\alpha_t^i dt + \sum_{j=1}^m \sigma_t^{ij} dW_t^j \right), \quad i = 1, \dots, n, \end{aligned} \right\}, t \in [0, T], \quad (2.1)$$

where the \mathbb{R}^n -valued process $\{\boldsymbol{\alpha}_t\}_{t \in [0, T]} = \{(\alpha_t^i)_{i=1, \dots, n}\}_{t \in [0, T]}$ is the *mean rate of return*, and $\{\boldsymbol{\sigma}_t\}_{t \in [0, T]} = \{(\sigma_t^{ij})_{i=1, \dots, n}^{j=1, \dots, m}\}_{t \in [0, T]} \in \mathbb{R}^{n \times m}$ is the *variance-covariance* process. In order for Equations (2.1) to admit unique strong solutions, we impose the following regularity conditions on the coefficient processes $\boldsymbol{\alpha}_t$ and $\boldsymbol{\sigma}_t$:

Assumption 2.1. All the components of the processes $\{\boldsymbol{\alpha}_t\}_{t \in [0, T]}$ and $\{\boldsymbol{\sigma}_t\}_{t \in [0, T]}$ are predictable, and

$$\sum_{i=1}^n \int_0^t |\alpha_u^i| du + \sum_{i=1}^n \sum_{j=1}^m \int_0^t \sigma_u^{ij2} du < \infty, \quad \text{for all } t \in [0, T], \mathbb{P}\text{-a.s.}$$

To ease the exposition, we introduce the following notation: for an integrable \mathbb{R}^m -valued process $\boldsymbol{\gamma}_t = (\gamma_t^i)_{i=1, \dots, m}$, and a sufficiently regular \mathbb{R}^m -valued process $\boldsymbol{\pi}_t =$

$(\pi_t^j)_{j=1,\dots,m}$ we write

$$\int_0^t \gamma_u du := \sum_{i=1}^n \int_0^t \gamma_u^i du, \quad \int_0^t \pi_t d\mathbf{W}_t := \sum_{j=1}^m \int_0^t \pi_u^j dW_u^j.$$

Further, we impose the following condition on the variance–covariance process σ_t :

Assumption 2.2. The matrix σ_t has, almost–surely, independent rows for all $t \in [0, T]$.

This assumption makes it impossible for different stocks to have the same diffusion structure. Otherwise, the market would either allow for arbitrage opportunities or redundant assets would exist. As a consequence of Assumption 2.2 we have that $n \leq m$ - the number of risky assets does not exceed the number of “sources of uncertainty”. Moreover, the inverse $(\sigma_t \sigma_t')^{-1}$ is easily seen to exist. The equation

$$\sigma_t \sigma_t' \zeta_t^M = \mu_t$$

uniquely defines a predictable stochastic process $\{\zeta_t^M\}_{t \in [0, T]}$, named the *Merton–proportion process*, where $\{\mu_t\}_{t \in [0, T]} = \{(\mu_t^i)_{i=1,\dots,n}\}_{t \in [0, T]}$ is the vector of *excess rates of return* with

$$\mu_t^i = \alpha_t^i - r$$

for $i = 1, \dots, n$. Note that the market model is incomplete whenever $n < m$. At this point we make another assumption on the market coefficients:

Assumption 2.3. The market coefficients $\{\alpha_t\}_{t \in [0, T]}$, $\{\sigma_t\}_{t \in [0, T]}$ and r , and the corresponding Merton–proportion process satisfy

$$\mathbb{E} \left[\exp \left(\int_0^T \|\zeta_t^M \sigma_u\|^2 du \right) \right] < \infty,$$

and the stochastic process $\sigma'(\sigma\sigma')^{-1}\sigma$ is uniformly bounded. In addition, we assume that there exists a constant $c > 0$ such that

$$\|\sigma'(\sigma\sigma')^{-1}\sigma_t \mu_t\| \leq c, \quad \forall t \in [0, T], \quad \mathbb{P} - a.s..$$

2.2. Trading strategies and wealth. Let \mathcal{P} denote the predictable σ -algebra on $[0, T] \times \Omega$. The control variables are the proportions of current wealth the investor invests in the assets. More precisely, we have the following formal definition:

Definition 2.4. An \mathbb{R}^n -valued stochastic process $\{\zeta_t\}_{t \in [0, T]} = \{(\zeta_t^i)_{i=1,\dots,n}\}_{t \in [0, T]}$ is called an *admissible portfolio–proportion process* if it is predictable (i.e. \mathcal{P} -measurable) and it satisfies

$$\mathbb{E} \left[\int_0^t |\zeta_u'(\alpha_u - r\mathbf{1})| du + \int_0^t \|\zeta_u' \sigma_u\|^2 du \right] < \infty, \quad \text{for all } t \in [0, T]. \quad (2.2)$$

Here ζ_t' denotes the transpose of ζ_t , $\mathbf{1} = (1, \dots, 1)'$ is a n -dimensional column vector all of whose coordinates are equal to 1, and $\|\mathbf{x}\|$ is the standard Euclidean norm. The set of admissible strategies will be denoted by \mathcal{A} .

Given a portfolio–proportion process ζ_t , we interpret its n coordinates as the proportions of the current wealth X_t^ζ invested in each of the n stocks. In order for the portfolio to be self–financing, the remaining wealth $X_t^\zeta(1 - \sum_{i=1}^n \zeta_t^i)$ is assumed to

be invested in the riskless bond S_t^0 . If this quantity is negative, we are effectively borrowing at the rate $r > 0$. No short-selling restrictions are imposed, hence the proportions ζ_t^i are allowed to be negative, and they are not a priori bounded. The equation governing the evolution of the total wealth $\{X_t^\zeta\}_{t \in [0, T]}$ of the investor using the portfolio-proportion process $\{\zeta_t\}_{t \in [0, T]}$ is given by

$$\begin{aligned} dX_t^\zeta &= X_t^\zeta \left(\zeta_t' \alpha_t dt + \zeta_t' \sigma_t d\mathbf{W}_t \right) + (1 - \zeta_t' \mathbf{1}) r X_t^\zeta dt \\ &= X_t^\zeta \left((r + \zeta_t' \mu_t) dt + \zeta_t' \sigma_t d\mathbf{W}_t \right), \end{aligned} \quad (2.3)$$

where we recall that $\mu_t^i = \alpha_t^i - r$ for $i = 1, \dots, n$. Under the regularity conditions imposed on ζ_t via Expression (2.2), Equation (2.3) admits a unique strong solution given by

$$X_t^\zeta = X(0) \exp \left\{ \int_0^t \left(r + \zeta_u' \mu_u - \frac{1}{2} \|\zeta_u' \sigma_u\|^2 \right) du + \int_0^t \zeta_u' \sigma_u d\mathbf{W}_u \right\}. \quad (2.4)$$

The initial wealth $X^\zeta(0) = X(0) \in (0, \infty)$ is considered to be exogenously given. As a consequence of Assumption 2.3, and using Expression (2.2), a strategy ζ is admissible if and only if it is a predictable process such that

$$\mathbb{E} \left[\int_0^T \|\zeta_u' \sigma_u\|^2 du \right] < \infty. \quad (2.5)$$

Indeed we have

$$\zeta_u' \mu_u = (\sigma_u' \zeta_u)' (\sigma_u' \zeta_u)^M \leq \|\zeta_u' \sigma_u\| \|\zeta_u^M \sigma_u\|,$$

by the Cauchy–Buniakowski–Schwarz inequality. Thus, Inequality (2.5) follows from Assumption 2.3, Expression (2.2) and the Cauchy–Buniakowski–Schwarz inequality.

Before proceeding any further, we briefly explain what is the stochastic control problem considered in this paper. Our aim is to find a strategy ζ^* such that its associated wealth process X^{ζ^*} maximizes the quantity

$$\mathbb{E}[U_p(X_T^\zeta)], \quad U_p(x) := \frac{x^p}{p}, \quad p < 1.$$

Any candidate solutions ζ must belong to the (time-dependent) set of risk-admissible strategies \mathcal{A}^p , which is determined via a dynamic risk constraint. We make precise the definition of the wealth process in Section 2.2 below, and we introduce the formal definition of the dynamic risk constraint imposed on the investor in Sections 2.3 and 2.4. The exact formulation of the maximization problem is finally given in Section 2.5.

It is clear from Expression (2.4) that the evolution of the wealth process X_t^ζ depends on the \mathbb{R}^n -dimensional process ζ_t only through two “sufficient statistics”, namely

$$\zeta_t^\mu := \zeta_t' \mu_t, \quad \text{and} \quad \zeta_t^\sigma := \|\zeta_t' \sigma_t\|. \quad (2.6)$$

These will be referred to in the sequel as *portfolio rate of return* and *portfolio volatility*, respectively. The expression appearing inside the first integral in Equation (2.4) will be given its own notation; namely, we define the quadratic function $\tilde{Q} : \mathbb{R}^2 \rightarrow \mathbb{R}$ as

$$\tilde{Q}(\zeta^\mu, \zeta^\sigma) := r + \zeta^\mu - \frac{1}{2} \zeta^\sigma^2. \quad (2.7)$$

It is also useful to define the random field $Q : \Omega \times [0, \infty) \times \mathbb{R}^n \rightarrow \mathbb{R}$

$$Q(t, \zeta) := \tilde{Q}(\zeta' \mu_t, \|\zeta' \sigma_t\|).$$

2.3. Projected distribution of wealth. For the purposes of risk measurement, it is common practice to use an approximation of the distribution of the investor’s wealth at a future date. Given the current time $t \geq 0$, and a length $\tau > 0$ of the measurement horizon $[t, t + \tau)$, the *projected distribution* of the wealth from trading is calculated under the simplifying assumptions that

- (1) the proportions of the wealth $\{\zeta_s\}_{s \in [t, t+\tau)}$ invested in various securities, as well as
- (2) the market coefficients $\{\alpha_s\}_{s \in [t, t+\tau)}$ and $\{\sigma_s\}_{s \in [t, t+\tau)}$

stay constant and equal to their present values throughout the time interval $[t, t + \tau)$. The wealth Equations (2.3) and (2.4) yield that the *projected wealth loss* is - conditionally on \mathcal{F}_t - distributed as $L = L(X_t, \zeta_t^\mu, \zeta_t^\sigma)$, where the law of $L(x, \zeta^\mu, \zeta^\sigma)$ is the one of

$$x \left(1 - \exp(Y(\zeta^\mu, \zeta^\sigma)) \right).$$

Here $Y(\zeta^\mu, \zeta^\sigma)$ is a normal random variable with mean $\tilde{Q}(\zeta^\mu, \zeta^\sigma)\tau$ and standard deviation $\sqrt{\tau}\zeta^\sigma$. The quantities ζ_t^μ and ζ_t^σ are the portfolio rate of return and volatility, defined in Equation (2.6). In the upcoming sections we turn our focus to risk measurements associated to the *relative projected wealth gain*, which will be defined as the distribution of the quantity

$$\frac{X^\zeta(t + \tau-) - X_t^\zeta}{X_t^\zeta}.$$

This is not a technical requirement, and the method developed in Sections 2.4 to 3 still holds for risk measurements in absolute terms. The economic implications, however, may be stark, and the definition of the risk constraints below would require a certain recursive structure. The latter in the sense that admissibility (risk-wise) at time t will depend on the choice of the strategy at all previous times. We elaborate further on this in Remark 2.6. The measurement horizon τ and the market coefficients will play the role of “global variables”.

2.4. The risk constraints. In this section we introduce the risk constraints that will be imposed on the trading strategies. We keep the presentation as general as possible and make only sufficient assumptions on the risk measures. These allow us to show existence (and in some cases uniqueness) of optimal, constrained trading strategies. We begin by making precise how the risk of a given strategy is measured.

Let us define the gain over time interval $[t, t + \tau]$ by $\Delta_\tau X_t^\zeta := X_{t+\tau-}^\zeta - X_t^\zeta$, and let $(\rho_t)_{t \in [0, T]}$ be a family of maps ρ_t with

$$\rho_t : \mathcal{C}_t \subset L^2(\mathcal{F}_T, \mathbb{P}) \rightarrow L^2(\mathcal{F}_t, \mathbb{P}),$$

where

$$\mathcal{C}_t := \left\{ \Delta_\tau X_t^\zeta / X_t^\zeta \mid \zeta \text{ is an admissible strategy} \right\}.$$

Notice that for all $t \in (0, T]$, we have that $\mathcal{C}_t \subset L^2(\mathcal{F}_T, \mathbb{P})$. We also define $\mathcal{C}_0 := L^2(\mathcal{F}_T, \mathbb{P})$. For a given admissible $(\tilde{\zeta}_s)_{s \in [0, t]}$ and $\zeta \in \mathbb{R}^n$ we define the strategy $\bar{\zeta} :$

$\Omega \times [0, t + \tau) \rightarrow \mathbb{R}^n$ as $\bar{\zeta}_s = \tilde{\zeta}_s$ for $s < t$ and $\bar{\zeta}_s = \zeta$ for $t \leq s < t + \tau$. By definition of the wealth process we obtain that $X_t^{\bar{\zeta}} = X_{t-}^{\tilde{\zeta}}$, moreover (under the assumptions made in Section 2.3) the quantity $\Delta_\tau X_t^{\bar{\zeta}}/X_t^{\bar{\zeta}}$ depends exclusively on ζ , and not on $\tilde{\zeta}$. In order to establish the risk constraints, we define the acceptance sets

$$\mathcal{A}_t^{\rho, \tilde{\zeta}}(\omega) := \left\{ \zeta \in \mathbb{R}^n \mid \rho_t \left(\frac{\Delta_\tau X_t^{\bar{\zeta}}}{X_t^{\bar{\zeta}}} \right) (\omega) \leq K_t(\omega) \right\} \quad t \in [0, T],$$

where K_t is a real-valued, exogenous, predictable process that satisfies $K_t \geq \rho_t(0)$ for all t in $[0, T]$, P -almost surely. Notice that $\zeta = 0$ is in the constraint set. We observe that, by construction, the sets $\mathcal{A}_t^{\rho, \tilde{\zeta}}$ are independent of $\tilde{\zeta}$, and we shall simply write \mathcal{A}_t^ρ . In analogous fashion we will slightly abuse notation and write $\Delta_\tau X_t^\zeta/X_t^\zeta$ for $\Delta_\tau X_t^{\bar{\zeta}}/X_t^{\bar{\zeta}}$. It follows from Equation (2.3) that in fact

$$\frac{\Delta_\tau X_t^\zeta}{X_t^\zeta} = \mathfrak{E}(\zeta, t) - 1,$$

where

$$\mathfrak{E}(\zeta, t) := \exp \left\{ \int_t^{t+\tau} \left(r + \zeta^\mu - \frac{1}{2} \zeta^\sigma \right) du + \int_t^{t+\tau} \zeta^\sigma \sigma_u dW_u \right\},$$

and ζ^μ and ζ^σ are defined in Expression (2.6). Hence, the expressions for the sets \mathcal{A}_t^ρ may be rewritten as

$$\mathcal{A}_t^\rho(\omega) = \left\{ \zeta \in \mathbb{R}^n \mid \rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega) \leq K_t(\omega) \right\}. \quad (2.8)$$

Moreover, under the assumption that μ , σ and ζ remain (for the purpose of risk assessment) constant over $[t, t + \tau)$, we may write

$$\mathfrak{E}(\zeta, t) := \exp \{ r\tau \} \cdot \exp \left\{ \tau \left(\zeta^\mu - \frac{1}{2} \zeta^\sigma \right) \right\} \cdot \exp \left\{ \zeta^\sigma \Delta_\tau W_t \right\},$$

and we shall denote by $\mathfrak{E}_1(\zeta, t)$ and $\mathfrak{E}_2(\zeta, t)$ the second and third factors of $\mathfrak{E}(\zeta, t)$, respectively.

We make the following assumption on the family $(\rho_t)_{t \in [0, T]}$:

Assumption 2.5. The family of maps

$$\rho_t : \mathcal{C}_t \subset L^2(\mathcal{F}_T, \mathbb{P}) \rightarrow L^2(\mathcal{F}_t, \mathbb{P})$$

satisfies that the mapping

$$(\zeta, (\omega, t)) \mapsto \rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega)$$

is a Carathéodory function; that is, for every (ω, t) in $\Omega \times [0, T]$, the map $\zeta \mapsto \rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega)$ is continuous and for every ζ in R^n the map $(\omega, t) \mapsto \rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega)$ is \mathcal{P} -measurable.

An example of a family $(\rho_t)_{t \in [0, T]}$ that satisfies Assumption 2.5 is the following: Let $l : \mathbb{R} \rightarrow \mathbb{R}$ be a convex, non-decreasing continuous and non-constant function¹ with $l(-\infty) < +\infty$. Assume that the filtration $\{\mathcal{F}_t\}_{t \in [0, T]}$ is generated by the Brownian

¹Such functions are usually referred to as “loss functionals”.

motion $\{\mathbf{W}_t\}_{t \in [0, T]}$ and that $\sigma_t^{i,j} := \sigma^{i,j}(t, W_t)$ and $\mu_t := \mu(t, W_t)$ where $\sigma^{i,j}$ and μ are deterministic Borelian functions. We set

$$\rho_t(-\mathfrak{E}(\zeta, t) - 1) = \mathbb{E}[l(\exp(r(\tau + \zeta x \frac{1}{2} \|\zeta y\|^2) + xy \Delta_\tau W_0))]_{x=\mu(t, W_t), y=\sigma(t, W_t)}$$

so that $\rho_t(-\mathfrak{E}(\zeta, t) - 1) = \mathbb{E}[l(-\mathfrak{E}(\zeta, t) - 1) | \mathcal{F}_t]$, \mathbb{P} -almost surely. Then the family $(\rho_t)_{t \in [0, T]}$ satisfies Assumption 2.5. Indeed, fix \mathbf{a} in \mathbb{R}^n and let ζ in \mathbb{R}^n . Then, by monotonicity of the exponential and l we have that:

$$l(-\infty) \leq l(-\mathfrak{E}(\zeta, t) + 1) \leq l(1).$$

Hence, Lebesgue's Dominated Convergence Theorem implies that:

$$\lim_{\zeta \rightarrow \mathbf{a}} \rho_t(-\mathfrak{E}(\zeta, t) - 1) = \rho_t(-\mathfrak{E}(\mathbf{a}, t) - 1), \quad \forall t \in [0, T].$$

Finally, since the filtration we consider is the Brownian filtration, the stochastic process $(\rho_t(\mathfrak{E}(\zeta, t) - 1))_{t \in [0, T]}$ is predictable.

Remark 2.6. If we were to consider risk constraints based not on the relative projected wealth loss, but only on the quantities $\Delta_\tau X_t^\zeta$, then the acceptance sets defined in Expression 2.8 would depend on $(\zeta_s)_{s \in [0, t]}$. More precisely, the set of risk-admissible strategies would be

$$\mathcal{A} := \left\{ \zeta = (\zeta_s)_{s \in [0, T]} \mid \zeta \text{ is admissible and } \zeta_t \in \mathcal{A}_t^{\rho, \zeta \mathbf{1}_{[0, t]}} \right\}$$

In the case where ρ_t is a \mathcal{F}_{t-} -coherent family, i.e. if $\rho_r(XY) = X \rho_t(Y)$ for all $X \in \mathcal{F}_{t-}$, then risk constraints in absolute terms are generated by inequalities of the form

$$X_t^\zeta \rho_t \left(\frac{\Delta_\tau X_t^\zeta}{X_t^\zeta} \right) \leq K_t.$$

This follows from the fact that the wealth level at time t is a \mathcal{F}_{t-} -measurable random variable. The structure then reverts to that of risk constraints in relative terms, except for a redefinition of the risk bound as $\tilde{K}_t(\omega) := K_t(\omega)/X_t(\omega)$. Notice that if $K_t \equiv K \in \mathbb{R}_+$, then \tilde{K}_t would be a decreasing function of wealth. In other words, highly capitalized investors would face more stringent constraints. This could lend an approach to dealing with the too-big-to-fail problem, and could be further tweaked by allowing K_t to depend on the state of nature. It is, however, beyond the scope of this paper to discuss such policy-making issues, and we shall stick to the relative-measures-of-risk framework.

Remark 2.7. Note that $(\rho_t)_{t \in [0, T]}$ is not *stricto sensu* a dynamic risk measure, since every ρ_t is a priori not defined on the whole space $L^2(\mathcal{F}_T, P)$. As we have seen in the previous lines, defining the risk of every random variable in $L^2(\mathcal{F}_T, P)$ is not relevant for us, since we only need to evaluate the risk of the very specific random variables $\frac{\Delta_\tau X_t^\zeta}{X_t^\zeta}$.

2.5. The optimization problem. We finish the section by formulating our central problem. Given a choice of a dynamic risk measure ρ satisfying Assumption 2.5 and a final date T , we are searching for a portfolio–proportion process $\zeta_t^* \in \mathcal{A}_t^\rho$ which maximizes the p –CRRA utility $U_p(x) = \frac{x^p}{p}$, $p < 1$, of the final wealth among all the portfolios satisfying the same constraint. In other words, for all $t \in [0, T]$ and $\zeta_t \in \mathcal{A}_t^\rho = \left\{ \zeta \in \mathbb{R}^n \mid \rho_t(\mathfrak{E}(\zeta, t) - 1) \leq K_t \right\}$

$$\mathbb{E}[U_p(X_T^{\zeta^*})] \geq \mathbb{E}[U_p(X_T^\zeta)]. \quad (2.9)$$

This problem has the following economic motivation: Risk managers limit the risk exposure of their traders by imposing risk constraints on their strategies. This can be regarded as an external risk management mechanism. In our model this is represented by the risk measures. On the other hand, traders have their own attitudes towards risk, which are reflected by the risk aversion of the CRRA utility. However, $p \in [0, 1)$ is known to reflect a risk seeking attitude of the trader. The risk manager cannot constrain the trader’s risk preferences. In order to deal with this, risk constraints on the trader’s strategies must be imposed.

3. Analysis

In this section we prove the existence of an optimal investment strategy. For simplicity we consider the case $p \in (0, 1)$ (analogous arguments apply with minor modifications to $p < 0$). In order to do so, we make use of the powerful theory of backward stochastic differential equations (BSDEs). Let

$$\mathcal{A}^\rho := \left\{ \zeta = (\zeta_t)_{t \in [0, T]} \in \mathcal{A} \mid \zeta_t \in \mathcal{A}_t^\rho, \forall t \in [0, T] \right\},$$

where \mathcal{A} is the set of admissible strategies in the sense of Definition 2.4, and \mathcal{A}_t^ρ is defined by (2.8). We recall that we consider the maximization problem

$$\max_{\zeta \in \mathcal{A}^\rho} \mathbb{E}[U_p(X_T^\zeta)].$$

By means of Equation (2.4) we may write

$$U_p(X_t^\zeta) = U_p(X(0)) \exp \left(\int_0^t p \tilde{Q}(\zeta_u^\mu, \zeta_u^\sigma) du + \int_0^t p \zeta_u' \sigma_u d\mathbf{W}_u \right),$$

where ζ_u^μ and ζ_u^σ are as in Expression (2.6), and \tilde{Q} is defined in (2.7). In analogous fashion as done in [13], let us introduce the auxiliary process

$$R_t^\zeta := U_p(X(0)) \exp \left(Y_t + \int_0^t p \tilde{Q}(\zeta_u^\mu, \zeta_u^\sigma) du + \int_0^t p \zeta_u' \sigma_u d\mathbf{W}_u \right),$$

where (Y, Z) is a solution to the BSDE

$$Y_t = 0 - \int_t^T Z_u d\mathbf{W}_u - \int_t^T h(u, Z_u) du, \quad t \in [0, T]. \quad (3.1)$$

The function $h(t, z)$ should be chosen in such a way that

- a) the process R^ζ is a supermartingale, $R_T^\zeta = U_p(X_T^\zeta)$ and there exists a constant $c_0 > 0$ (which does not depend on ζ), such that $R^\zeta(0) = c_0$ for every $\zeta \in \mathcal{A}^\rho$,
- b) there exists at least one element ζ^* in \mathcal{A}^ρ such that R^{ζ^*} is a martingale.

We shall verify ex-post that the function $h(t, z)$ in question satisfies the measurability and growth conditions required to guarantee existence of solutions to Equation (3.1).

Before going further we explain why achieving this would provide a solution to Problem (2.9): If we were able to construct such a family of processes R^ζ , then we would obtain that ζ^* is an optimal strategy for Problem (2.9) with initial capital $X(0) > 0$ independent of ζ . Indeed let ζ any element of \mathcal{A}^ρ , then using (a) and (b) we have

$$\mathbb{E}(U_p(X_T^\zeta)) = \mathbb{E}(R_T^\zeta) \leq R^\zeta(0) = R^{\zeta^*}(0) = \mathbb{E}(R_T^{\zeta^*}).$$

This method is known as the *martingale optimality principle*.

Let us now perform a multiplicative decomposition of R^ζ into martingale and an increasing process. Given a continuous process M , we denote by $\mathcal{E}(M)$ its stochastic exponential:

$$\mathcal{E}(M_t) := \exp\left(M_t - \frac{1}{2}\langle M \rangle_t\right),$$

where $\langle M \rangle$ denotes the quadratic variation. Then

$$R_t^\zeta = \frac{(X(0))^p}{p} \mathcal{E}\left(\int_0^t (p \zeta'_u \sigma_u + Z_u) d\mathbf{W}_u\right) \exp\left(\int_0^t g(u, Z_u) du\right), \quad (3.2)$$

where

$$g(u, z) := h(u, z) + \frac{1}{2}\|z\|^2 + pr + p\zeta'_u(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z) + \frac{p^2 - p}{2}\|\zeta'_u \boldsymbol{\sigma}_u\|^2.$$

Since R^{ζ^*} should be a supermartingale for every admissible ζ_u (and a martingale for some element ζ_u^*), then g has to be a non-positive process. With this in mind, a suitable candidate would be

$$h(u, z) := -pr - \frac{1}{2}\|z\|^2 + \inf_{\zeta_u \in \mathcal{A}_u} \left\{ -p\zeta'_u(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z) + \frac{p - p^2}{2}\|\zeta'_u \boldsymbol{\sigma}_u\|^2 \right\},$$

which leads to

$$\begin{aligned} h(u, z) &= -pr - \frac{1}{2}\|z\|^2 + \frac{p}{2(p-1)}\|\boldsymbol{\sigma}'_u(\boldsymbol{\sigma}\boldsymbol{\sigma}')^{-1}(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z)\|^2 \\ &\quad + \frac{p(1-p)}{2} \text{dist}\left(\frac{\boldsymbol{\sigma}'_u(\boldsymbol{\sigma}\boldsymbol{\sigma}')^{-1}(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z)}{1-p}; \mathcal{A}_u^\rho \boldsymbol{\sigma}_u\right)^2. \end{aligned} \quad (3.3)$$

If in addition we let

$$\tilde{z} := \frac{\boldsymbol{\sigma}'_u(\boldsymbol{\sigma}\boldsymbol{\sigma}')^{-1}(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z)}{1-p} \quad \text{and} \quad \tilde{\mathcal{A}}_u^\rho := \mathcal{A}_u^\rho \boldsymbol{\sigma}_u, \quad (3.4)$$

then

$$\text{dist}\left(\frac{\boldsymbol{\sigma}'_u(\boldsymbol{\sigma}\boldsymbol{\sigma}')^{-1}(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z)}{1-p}; \mathcal{A}_u^\rho \boldsymbol{\sigma}_u\right)^2 = \left\| \frac{\boldsymbol{\sigma}'_u(\boldsymbol{\sigma}\boldsymbol{\sigma}')^{-1}(\boldsymbol{\mu}_u + \boldsymbol{\sigma}_u z)}{1-p} - \zeta^{*\prime}_u \boldsymbol{\sigma}_u \right\|^2$$

with

$$\zeta^{*\prime}_u \boldsymbol{\sigma}_u \in \text{Proj}(\tilde{Z}_u, \tilde{\mathcal{A}}_u^\rho).$$

The available results on existence of solutions to BSDEs require, to begin with, the predictability of the driver h . In our case this is closely related to the predictability of ζ^* , in other words, to whether or not the candidate for an optimal strategy is acceptable.

Theorem 3.1. *Let Z be a predictable process such that*

$$\mathbb{E} \left(\int_0^T \|Z_u\|^2 du \right) < \infty.$$

Then for $(t, \omega) \in [0, T] \times \Omega$, the mapping

$$(t, \omega) \mapsto \text{dist}(\tilde{Z}_t(\omega), \tilde{\mathcal{A}}_t^\rho(\omega)),$$

where \tilde{Z} is as in Equation (3.4), is predictable. In addition there exists a predictable process ζ^ in \mathbb{R}^n such that*

$$\mathbb{E} \left(\int_0^T \|\zeta_u^{*'} \sigma_u\|^2 du \right) < \infty$$

and

$$\text{dist}(\tilde{Z}_t, \tilde{\mathcal{A}}_t^\rho) = \text{dist}(\tilde{Z}_t, \zeta_t^{*'} \sigma_t), \quad \forall t \in [0, T], P - a.s..$$

Proof. Let us define for $k \in \mathbb{N}$

$$\mathcal{A}_{t,k}^\rho(\omega) := \left\{ \zeta \in [-k, k]^n \mid \rho_t(\mathfrak{E}(\zeta, t))(\omega) - K_t(\omega) \leq 0 \right\}.$$

The purpose of artificially bounding the values of \mathcal{A}^ρ is to make use of the theory of compact-valued correspondences (see Appendix A). It follows from Lemma A.1 that for all $k \in \mathbb{N}$ and for all (t, ω) , the set $\mathcal{A}_{t,k}^\rho(\omega)$ is non-empty and compact. Moreover, Proposition A.3 guarantees that for all $t \in [0, T]$ and $k \in \mathbb{N}$, the correspondence $(\omega, t) \mapsto \tilde{\mathcal{A}}_{t,k}^\rho(\omega)$ is weakly \mathcal{P} -measurable (see Definition A.2 in the Appendix for the definition of weak measurability). Let $(C(\mathbb{R}^m), \mathcal{H})$ denote the space of non-empty, compact subsets of \mathbb{R}^m , equipped with the Hausdorff metric. This is a complete, separable metric space, in which $\tilde{\mathcal{A}}_{t,k}^\rho(\cdot)$ takes its values. Theorem A.4 then states that for $z \in \mathbb{R}^m$ and $t \in [0, T]$, the distance mapping

$$\delta(\omega, z) = \text{dist}(z, \mathcal{A}_{t,k}^\rho(\omega) \sigma_t)$$

is a Carathéodory one. Since the process \tilde{Z}_t is predictable and $z \mapsto \delta(z, \omega)$ is continuous for all $\omega \in \Omega$, the map

$$(\omega, t) \mapsto \text{dist}(\tilde{Z}_t(\omega), \mathcal{A}_{t,k}^\rho(\omega) \sigma_t)$$

is \mathcal{P} -measurable. Finally

$$\text{dist}(\tilde{Z}_t(\omega), \tilde{\mathcal{A}}_t^\rho(\omega)) = \inf_{k \in \mathbb{N}} \left\{ \text{dist}(\tilde{Z}_t(\omega), \mathcal{A}_{t,k}^\rho(\omega) \sigma_t) \right\},$$

thus the mapping $\omega \mapsto \text{dist}(\tilde{Z}_t(\omega), \tilde{\mathcal{A}}_t^\rho(\omega))$ is predictable as the pointwise infimum of predictable ones. We now turn our attention to the second claim. First we observe that since $\tilde{\mathcal{A}}_t^\rho(\omega)$ is closed (and contained in \mathbb{R}^m), the set

$$\bar{\mathcal{A}}_t^\rho(\omega) := \text{argmin}_{a \in \tilde{\mathcal{A}}_t^\rho(\omega)} \left\{ \text{dist}(\tilde{Z}_t(\omega), a) \right\}$$

is compact. It follows from the Measurable Maximum Theorem ([1], page 605) that the correspondence $(t, \omega) \mapsto \bar{\mathcal{A}}_t^\rho(\omega)$ is weakly \mathcal{P} -measurable. It is then implied by the Kuratowski–Ryll–Nardzewski Selection Theorem that $\bar{\mathcal{A}}^\rho(\cdot)$ admits a measurable

selection $\zeta^{*'}\sigma$; in other words, there exists a predictable process $\zeta^* : [0, T] \times \Omega \rightarrow \mathbb{R}^n$ such that

$$\text{dist}(\tilde{Z}_t(\omega), \tilde{\mathcal{A}}_t^\rho(\omega)) = \text{dist}(\tilde{Z}_t(\omega), \zeta^{*'}(t, \omega)\sigma(t, \omega)), \quad \zeta^{*'}(t, \omega)\sigma(t, \omega) \in \tilde{\mathcal{A}}_t^\rho(\omega).$$

Finally using the fact that the strategy $(0, \dots, 0)$ belongs to $\tilde{\mathcal{A}}^\rho$ we have that

$$\begin{aligned} \int_0^T \|\zeta^{*'}_u \sigma_u\|^2 du &\leq 2 \int_0^T \|\zeta^{*'}_u \sigma_u - \tilde{Z}_u\|^2 du + 2 \int_0^T \|\tilde{Z}_u\|^2 du \\ &= 2 \int_0^T \text{dist}(\tilde{Z}_u, \tilde{\mathcal{A}}_u^\rho)^2 du + 2 \int_0^T \|\tilde{Z}_u\|^2 du \\ &\leq 4 \int_0^T \|\tilde{Z}_u\|^2 du < \infty. \end{aligned} \tag{3.5}$$

□

To finalize, we must show that the quadratic-growth BSDE (3.1) admits a solution. In the following we will make use of the notion of BMO-martingale.

Definition 3.2. A continuous martingale M is a BMO-martingale if there exists a positive constant $a > 0$ such that for every stopping time $\nu \leq T$,

$$\mathbb{E}[\langle M \rangle_T - \langle M \rangle_\nu | \mathcal{F}_\nu] \leq a, \quad \mathbb{P} - a.s..$$

We will use the following property of BMO-martingales (which can be found in [15]): if M is a BMO-martingale then $\mathcal{E}(M)$ is a true martingale.

We require the following result of Morlais [19, Theorem 2.5, Theorem 2.6 and Lemma 3.1], which extends the results of Kobylanski [17]:

Theorem 3.3. Let $h : [0, T] \times \Omega \times \mathbb{R}^m \rightarrow \mathbb{R}$ be measurable. Assume that there exist a predictable process α and positive constants C_1, C_2 satisfying $\alpha \geq 0$ and

$$\int_0^T \alpha_s ds \leq C_1, \quad \mathbb{P} - a.s..$$

If h is such that

- (1) $z \mapsto h(u, z)$ is continuous
- (2) $|h(u, z_1) - h(u, z_2)| \leq C_2(\alpha_u + \|z_1\| + \|z_2\|)(\|z_1 - z_2\|), \quad \forall z_1, z_2,$ and $|h(u, 0)| \leq \alpha_u.$

then the BSDE (3.1) with driver h admits a unique solution (Y, Z) , where Y and Z are predictable processes with Y bounded and Z satisfying $\mathbb{E} \left(\int_0^T \|Z_t\|^2 dt \right)^{\frac{1}{2}} < \infty$. In addition, the process $\int_0^\cdot Z_s d\mathbf{W}_s$ is a BMO martingale and hence $\mathcal{E} \left(\int_0^\cdot Z_s d\mathbf{W}_s \right)$ is a true martingale.

The previous result allows us to show that the BSDE (3.1) with driver given by Equation (3.3) admits a unique solution. Note that the fact that $\mathcal{E} \left(\int_0^\cdot Z_s d\mathbf{W}_s \right)$ is a true martingale is essential in our approach, since it implies that the process R^{ζ^*} is a (true) martingale for some element ζ^* .

Corollary 3.4. *There exists a unique pair of predictable processes (Y, Z) , with Y bounded and Z satisfying $\mathbb{E} \left(\int_0^T \|Z_t\|^2 dt \right) < \infty$, solution to the BSDE (3.1) with driver given by Equation (3.3). In addition, the processes $\int_0^\cdot Z_s d\mathbf{W}_s$ and $\int_0^\cdot \zeta_u^{*'} \sigma_u d\mathbf{W}_u$ are BMO–martingales with ζ^* given by Theorem 3.1.*

Proof. We combine the previous results to apply Theorem 3.3. First, measurability of h is guaranteed by Theorem 3.1. Then, the continuity in z of the driver is straightforward, as are the growth conditions, given Assumption 2.3. Hence by Theorem 3.3, there exists a unique solution (Y, Z) to (3.1) such that Y is bounded, and $\int_0^\cdot Z_s d\mathbf{W}_s$ is a BMO–martingale. By definition, the latter means that there exists a positive constant $a > 0$ such that for every stopping time ν ,

$$\mathbb{E} \left[\int_\nu^T \|Z_s\|^2 ds \middle| \mathcal{F}_\nu \right] \leq a, \mathbb{P} - a.s..$$

Thus, by Estimate (3.5), we have for any stopping time ν that

$$\mathbb{E} \left[\int_\nu^T \|\zeta_u^{*'} \sigma_u\|^2 du \middle| \mathcal{F}_\nu \right] \leq 4\mathbb{E} \left[\int_\nu^T \|\tilde{Z}_u\|^2 du \middle| \mathcal{F}_\nu \right],$$

which shows that $\int_0^\cdot \zeta_u^{*'} \sigma_u d\mathbf{W}_u$ is a BMO–martingale since $\sigma'(\sigma\sigma')^{-1}\sigma\mu$ is uniformly bounded by Assumption 2.3. \square

We conclude with the existence of an optimal strategy to Problem (2.9).

Theorem 3.5. *Under the assumptions made above there exists an acceptable strategy ζ^* that solves Problem (2.9). If we define the value function $v(x)$ as:*

$$v(x) := \max_{\zeta \in \mathcal{A}^p} \mathbb{E}(U_p(X_T^\zeta)), \quad x > 0$$

with \mathcal{A}^p the set of admissible \mathbb{R}^n -valued predictable processes ζ such that $\zeta_t \in \mathcal{A}_t^p$ for all t in $[0, T]$ and $X^\zeta(0) = x$, then it holds that

$$v(x) = U_p(x) \exp(Y_0).$$

Here (Y, Z) is a solution to the BSDE (3.1) with driver given by Equation (3.3) and

$$\zeta_u^{*'} \sigma_u \in \text{Proj}(\tilde{Z}_u, \tilde{\mathcal{A}}_u^p).$$

Proof. The existence of a solution to the BSDE (3.1) is guaranteed by Corollary 3.4. Furthermore, by Corollary 3.4 the process

$$\mathcal{E} \left(\int_0^t (p \zeta_u^{*'} \sigma_u + Z_u) d\mathbf{W}_u \right)$$

is a true martingale since $\int_0^\cdot Z_s d\mathbf{W}_s$ and $\int_0^\cdot \zeta_u^{*'} \sigma_u d\mathbf{W}_u$ are BMO–martingales (with ζ^* given as in Theorem 3.1). Now, as in [13, Theorem 14], for any admissible ζ' , the process $R^{\zeta'}$ given by Equation (3.2) is a supermartingale. Indeed, by construction g is non-positive and the stochastic exponential $\mathcal{E} \left(\int_0^t (p \zeta_u' \sigma_u + Z_u) d\mathbf{W}_u \right)$ is local martingale. Let $(\tau_n)_n$ be a localizing sequence associated to it. We have for every

n (and $s \leq t$) that: $\mathbb{E}[R_{t \wedge \tau_n}^{\zeta'} | \mathcal{F}_s] \leq R_{s \wedge \tau_n}^{\zeta'}$ and $R^{\zeta'}$ is a non-negative process. Thus, Fatou's Lemma implies that

$$\mathbb{E}[R_t^{\zeta'} | \mathcal{F}_s] \leq R_s^{\zeta'}.$$

Using the martingale optimality principle, we have that the processes R^ζ are well-defined and satisfy requirements (a) and (b). In addition, by construction, the processes ζ^* such that R^{ζ^*} is a martingale are those such that $\zeta_u^{*'} \sigma_u \in \text{Proj}(\tilde{Z}_u, \tilde{A}_u^\rho)$. Theorem 3.1 yields that these elements ζ^* are admissible strategies, thus optimal. Take such an optimal strategy ζ^* . We have that

$$v(x) = \mathbb{E}(U_p(X_T^{\zeta^*})) = \mathbb{E}(U_p(R_T^{\zeta^*})) = R^{\zeta^*}(0) = U_p(x) \exp(Y_0).$$

□

The previous result admits a dynamic version:

Theorem 3.6. *Let $v(t, x)$ be the dynamic value function defined as:*

$$v(t, x) := \text{esssup}_{\zeta \in \mathcal{A}^t} \mathbb{E} \left(U_p \left(x + \int_t^T \zeta_s X_s^\zeta \frac{dS_s}{S_s} \right) \middle| \mathcal{F}_t \right) \quad t \in [0, T], \quad x > 0,$$

where $\mathcal{A}^t := \{\zeta \in \mathcal{A}^\rho, \zeta_s = 0, s < t\}$. Then

$$v(t, x) = U_p(x) \exp(Y_t),$$

where (Y, Z) is a solution to the BSDE (3.1) with driver given by Equation (3.3) and

$$\zeta_u^{*'} \sigma_u \in \text{Proj}(\tilde{Z}_u, \tilde{A}_u^\rho).$$

Proof. Let ζ be any element of \mathcal{A} and let ζ^* be such that the associated R^{ζ^*} is a martingale. Then by definition of the R^ζ processes, we have that $R_t^\zeta = U_p(x) \exp(Y_t)$ since $\zeta_s = 0$ for $s < t$ and so

$$\begin{aligned} & \mathbb{E} \left(U_p \left(x + \int_t^T \zeta_s X_s^\zeta \frac{dS_s}{S_s} \right) \middle| \mathcal{F}_t \right) \\ &= \mathbb{E} \left(R_T^\zeta \middle| \mathcal{F}_t \right) \\ &\leq R_t^\zeta = U_p(x) \exp(Y_t) = \mathbb{E} \left(R_T^{\zeta^*} \middle| \mathcal{F}_t \right) = \mathbb{E} \left(U_p \left(x + \int_t^T \zeta_s^* X_s^{\zeta^*} \frac{dS_s}{S_s} \right) \middle| \mathcal{F}_t \right). \end{aligned}$$

Hence, $v(t, x) = U_p(x) \exp(Y_t)$. □

Remark 3.7. Sometimes one might be interested in another version of the dynamic value function above. Given an element ζ in \mathcal{A}^ρ they may consider the quantity

$$v(t, X_t^\zeta) := \text{esssup}_{\tilde{\zeta} \in \mathcal{A}^{t, \zeta}} \mathbb{E} \left(U_p \left(X_t^\zeta + \int_t^T \tilde{\zeta}_s X_s^{\tilde{\zeta}} \frac{dS_s}{S_s} \right) \middle| \mathcal{F}_t \right), \quad t \in [0, T],$$

where $\mathcal{A}^{t, \zeta} := \{\tilde{\zeta} \in \mathcal{A}^\rho, \tilde{\zeta}_s = \zeta_s, s \leq t\}$. Then we have that $v(t, X_t^\zeta) = U_p(X_t^\zeta) \exp(Y_t)$ where (Y, Z) is the unique solution of the BSDE (3.1) with driver given by Equation (3.3).

Remark 3.8. The stochastic process $\exp(Y_t)$ in the expression of the value function is sometimes called the *opportunity process*, since it gives the value of the optimal wealth with initial capital one unit of currency (see [20]).

Remark 3.9. Notice that for the sake of the explanation, we have chosen to fix the risk aversion coefficient p in $(0, 1)$ but we can also consider the case where $p < 0$. Then the driver h given by Equation (3.3) has to be modified suitably.

4. Time Consistent Distortion Risk Measures

In this section we define a broad class of families of risk measures that are *time consistent*. We show that, under the constraints imposed by members of this class, optimal investment strategies follow a three-fund separation behavior. Let

$$\rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega) := \exp\{r\tau\} \mathfrak{E}_1(\zeta, t)(\omega) \rho_0(\mathfrak{E}_2(\zeta, t)(\omega) - 1),$$

where

$$\mathfrak{E}_1(\zeta, t)(\omega) := \exp\left\{\tau\left(\zeta x - \frac{1}{2}\|\zeta y\|^2\right)\right\}\Big|_{x=\mu_t(\omega), y=\sigma_t(\omega)}$$

and

$$\rho_0(\mathfrak{E}_2(\zeta, t)(\omega) - 1) := \rho_0\left(\exp\{xy\Delta_\tau W_0\} - 1\right)\Big|_{x=\zeta_t(\omega), y=\sigma_t(\omega)}.$$

Here ρ_0 is a distortion risk measure, i.e.

$$\rho_0(X) = \int_{[0,1]} F_u^{X^{-1}} dD_u,$$

where $F^{X^{-1}}$ is the inverse CDF of X , and D is a distortion; i.e., it is right-continuous, increasing on $[0, 1]$, $D(0) = 0$ and $D(1) = 1$. The choice $D_u = \mathbf{1}_{\{u \geq 1-\alpha\}}$ yields VaR_α and $D_u = \frac{1}{\alpha}[u - (1 - \alpha)]^+$ yields TVaR_α . LEL_α can be recovered by choosing $D_u = \frac{1}{\alpha}[u - (1 - \alpha)]^+ \mathbf{1}_{\mu=0}$ (since LEL_α is TVaR_α computed under one of the risk neutral probability measures). Distortion risk measures form a rich class, which contains, among others: proportional hazards, proportional odds, Gaussian distortion and positive Poisson mixture (for more about these risk measures see Examples 3.5, 3.6, 3.7, 3.8 in [25]). It follows from direct computations that

$$\rho_t(\mathfrak{E}(\zeta, t) - 1) = \int_{[0,1]} \left[1 - \exp\left(\tilde{Q}(\zeta_t^\mu, \zeta_t^\sigma)\tau + N_u^{-1}\zeta_t^\sigma\sqrt{\tau}\right)\right] dD_u,$$

where N^{-1} denotes the inverse of the distribution function of the standard Gaussian law. Another example is given by considering $D_u = u$. In that case, it holds that $\rho_0(X) = \int_{[0,1]} x dF^X(x) = \mathbb{E}[X]$. In the light of this, one can see that Assumption (2.5) holds true. From this point on we work under the assumption that $K_t < 1$. This implies (quite naturally) that the risk should be smaller than the current position.

4.1. A common form of the risk constraints. Below we present some properties of the constraint sets \mathcal{A}^ρ .

Proposition 4.1. *Each constraint set \mathcal{A}_t^ρ can be expressed as*

$$\mathcal{A}_t^\rho = \{\zeta \in \mathbb{R}^m : f(\zeta' \mu_t, \|\zeta' \sigma_t\|) \leq K_t\},$$

for some function $f : \mathbb{R} \times [0, \infty) \rightarrow \mathbb{R} \cup \{\infty\}$, which satisfies

$$f \in C^1(\mathbb{R} \times [0, \infty)), \quad f(0, 0) \leq 0, \quad \lim_{\zeta \rightarrow \infty} f(\zeta' \mu_t, \|\zeta' \sigma_t\|) = 1, \quad \mathbb{P} \otimes dt - a.e..$$

Proof. The function f is defined by

$$f(x, y) = \int_{[0,1]} \left[1 - \exp \left((r + x - \frac{y^2}{2})\tau + N_u^{-1} y \sqrt{\tau} \right) \right] dD_u,$$

so it follows that $f \in C^1(\mathbb{R} \times [0, \infty))$, $f(0, 0) \leq 0$. Since

$$\lim_{\zeta \rightarrow \infty} \left[1 - \exp \left(\tilde{Q}(\zeta' \mu_t, \|\zeta' \sigma_t\|)\tau + N_u^{-1} \zeta_t^\sigma \sqrt{\tau} \right) \right] = 1, \quad \mathbb{P} \otimes dt - a.e.$$

then $\lim_{\zeta \rightarrow \infty} f(\zeta' \mu_t, \|\zeta' \sigma_t\|) = 1$. \square

The choice of the threshold $K_t < 1$ and Proposition 4.1 yield the compactness of the constraint sets associated with the risk measures considered in this section.

4.2. A three-fund separation result. In this section we further characterize the optimal investment strategy. Let us recall that ζ^* is given by

$$\zeta_u^{*'} \sigma_u \in \text{Proj}(\tilde{Z}_u, \tilde{\mathcal{A}}_u^\rho), \quad u \in [0, T].$$

Compactness of \mathcal{A}^ρ leads to compactness of $\tilde{\mathcal{A}}^\rho$ which in turn yields the existence of the projection.

Theorem 4.2. *There exist two stochastic processes β^{1*} and β^{2*} such that optimal strategy ζ^* can be decomposed as*

$$\zeta_t^* = \frac{\beta^{1*}}{1-p} \zeta_t^M + \beta_t^{2*} (\sigma_t \sigma_t')^{-1} \sigma_t Z_t, \quad 0 \leq t \leq T, \quad (4.1)$$

where $Z_t, 0 \leq t \leq T$ is part of the (Y, Z) solution of BSDE (3.1) with driver (3.3).

Proof. We cover the case $p \geq 0$ only (the $p < 0$ case can be obtained by an analogous argument). Recall that for a fixed path ω , the optimal strategy ζ_t^* solves

$$\zeta_t^* = \arg \min_{\zeta \in \mathcal{A}_t} \left\{ -p \zeta' (\mu_t + p \sigma_t Z_t) + \frac{p-p^2}{2} \|\zeta' \sigma_t\|^2 \right\}.$$

The convex, quadratic functional

$$\zeta \rightarrow H(t, \zeta) := -p \zeta' (\mu_t + p \sigma_t Z_t) + \frac{p-p^2}{2} \|\zeta' \sigma_t\|^2$$

is minimized over the constraint set \mathcal{A}_t at a point ζ_t^* , which is either an absolute minimum or else should be on the boundary of \mathcal{A}_t . Thus, for a fixed path, ζ_t^* minimizes $H(t, \zeta)$ over the constraint $f(\zeta' \mu_t, \|\zeta' \sigma_t\|) \leq K_t$. The solution ζ_t^* is not the zero vector, since the latter is not an absolute minimum and $f(0, 0) \leq 0$. For $\zeta \neq 0$, it follows that

$$\nabla f(\zeta' \mu_t, \|\zeta' \sigma_t\|) = f_1(\zeta' \mu_t, \|\zeta' \sigma_t\|) \mu_t - \frac{f_2(\zeta' \mu_t, \|\zeta' \sigma_t\|)}{\|\zeta' \sigma_t\|} \sigma_t \sigma_t' \zeta,$$

where f_1 and f_2 stand for the partial derivatives of the function f . According to the Karush–Kuhn Tucker Theorem, either $\nabla f(\zeta' \mu_t, \|\zeta' \sigma_t\|) = 0$ or else there is a positive λ such that

$$\nabla H(t, \zeta) = \lambda \nabla f(\zeta' \mu_t, \|\zeta' \sigma_t\|). \quad (4.2)$$

In both cases, straightforward computations show that ζ_t^* should have the form given in Equation (4.1). \square

Theorem 4.2 is a three-fund separation result. It states that a utility-maximizing investor who is subject to regulatory constraints will invest his wealth into three-funds: 1. the savings account; 2. a risky fund with return $\zeta_t^M, t \in [0, T]$; 3. a risky fund with return $(\sigma_t \sigma_t')^{-1} \sigma_t' Z_t, t \in [0, T]$. Most of the results in the financial literature are two-funds separation ones (optimal wealth being invested into a savings account and a risky fund). We would obtain such a two-funds separation result if we restricted our model to one in which stocks returns and volatilities were deterministic. It is a consequence of the randomness of the stocks returns and volatilities that the optimal investment includes an extra risky fund. Investments in the latter can be regarded as a hedges against the risk implied by stochastic stock returns and volatilities.

Remark 4.3. For the special case of $\rho_0 = \text{TVaR}_\alpha$ the associated acceptance set \mathcal{A}^ρ is convex; this is also the case when $\rho_0 = \text{VaR}_\alpha$, whenever $\alpha \in [0, 0.5]$. The convexity of \mathcal{A}^ρ implies the uniqueness of optimal trading strategy ζ^* , a fact that turns out to be useful in numerical implementations.

5. A Numerically Implemented Example

In this section we present numerical simulations for the constrained optimal strategies and the associated constrained *opportunity processes*. Recall that by opportunity process we mean the process $\exp(Y_t)$, which appears in the value function $v(t, x)$ in Theorem 3.6; that is $v(t, x) = \frac{x^p}{p} \exp(Y_t)$. The opportunity process represents the value function of an investor with initial capital one dollar. It is a stochastic process and in the figures below we present one sample path. For simplicity and the numerical tractability of the analysis we assume that we deal with one risky asset ($n = 1$), one bond with rate zero ($r = 0$) and one Brownian motion ($m = 1$). In addition, we assume that the risky asset is given by the following SDE:

$$dS_t = S_t(\mathbf{1}_{[-1,1]}(W_t)dt + dW_t), \quad t \in [0, 1] \quad (T = 1), \quad S^0 = 1.$$

Our simulations require the use of numerical schemes for quadratic growth BSDEs. We use the scheme of Dos Reis and Imkeller [14, 9]. The latter relies, in a nutshell, on a truncation argument of the driver, and it reduces the numerical-simulation problem to one of a BSDE with a Lipschitz-growth driver. Here we use the so-called *forward scheme* of Bender and Denk [4].

In Figure 1 we illustrate the opportunity processes arising from imposing VaR, TVar and LEL. We have used the following set of parameters: $p=0.85$, $\alpha=0.10$, $K=0.3$ and $T = 1$. The time discretization is $1/15$ and $\tau=1/15$. The unconstrained opportunity process is also presented. The corresponding trading strategies are shown in Figure 2. We observe a spike in the opportunity process that may be explained by gambling; indeed looking at the TVaR constrained optimal strategy we see that it differs considerably from the unconstrained one (in which the stock is shorted). This finding supports the idea that risk constraints reduce speculation.

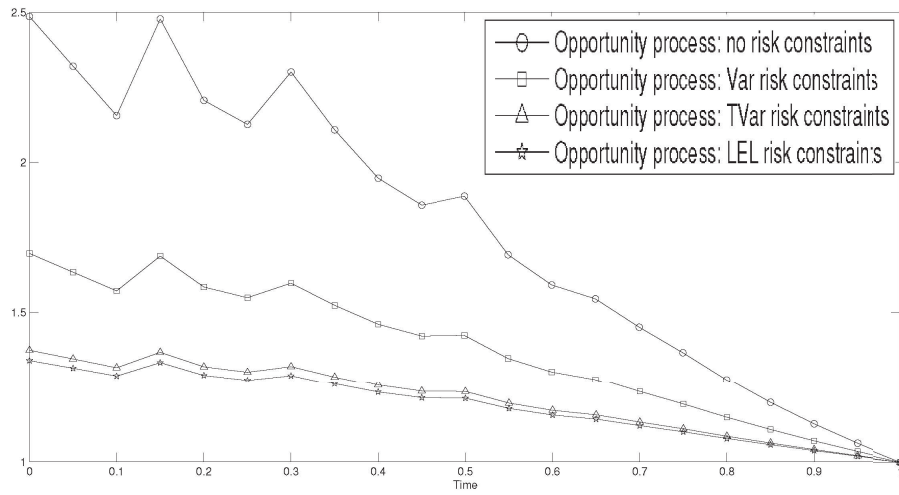


FIGURE 1. Constrained and unconstrained opportunity processes.

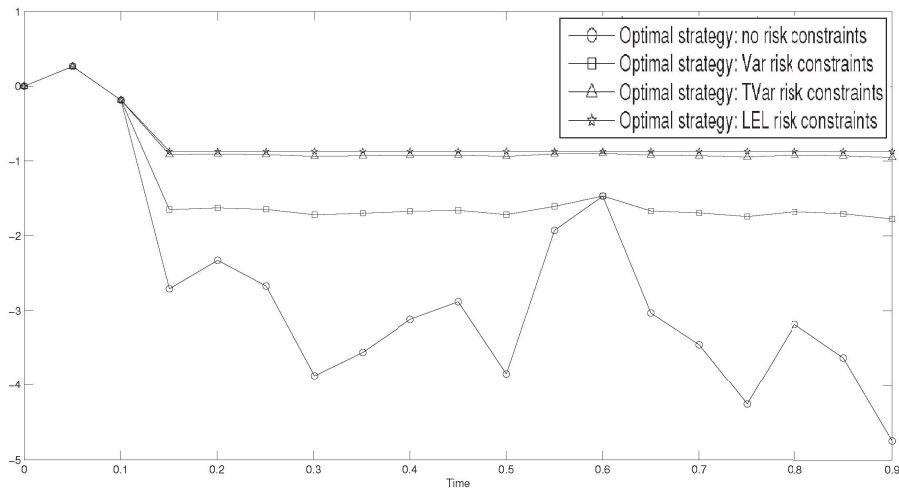


FIGURE 2. Constrained and unconstrained optimal strategies.

6. Conclusions

We have analyzed, within an incomplete-market framework, the portfolio-choice problem of a risk averse agent (who is characterized by CRRA preferences), when risk constraints are imposed continuously throughout the investment phase. Using BSDE technology, in the spirit of [13], has allowed us to consider a broad range of risk measures that give rise to the risk constraints, the latter being (possibly) time-dependent. In order to use such technology, we have made use of Measurable Selections

theory, specifically when addressing the issue of measurability of the driver of the BSDE at hand. We have characterized the optimal (constrained) investment strategies, and in the case of distortion risk measures we have provided explicit expressions for them. Here we have shown that optimal strategies may be described as investments in three funds, which is in contrast with the classical two-fund separation theorems. Finally, using recent results in [14], we have provided some examples that showcase the way in which our dynamic risk constraints limit investment strategies and impact utility at maturity.

Appendix A. Properties of the Constraint Sets \mathcal{A}_t^ρ

Several analytical properties of the (instantaneous) constraint sets \mathcal{A}_t^ρ are established in this section. The analysis requires some core concepts of the theory of measurable correspondences². We require the following auxiliary correspondences:

$$\mathcal{A}_{t,k}^\rho(\omega) := \left\{ \zeta \in [-k, k]^n \mid \rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega) - K_t(\omega) \leq 0 \right\}, \quad k \in \mathbb{N}.$$

The purpose of artificially bounding the values of \mathcal{A}^ρ is to make use of the theory of compact-valued correspondences, which exhibit many desirable properties.

Lemma A.1. *For any $m \in \mathbb{N}$, the correspondence $\mathcal{A}_{t,k}^\rho : [0, T] \times \Omega \rightarrow \mathbb{R}^n$ is non-empty and compact valued for almost all $(t, \omega) \in [0, T] \times \Omega$.*

Proof. The non-vacuity follows from the fact that $\zeta \equiv 0$, i.e. no wealth invested in risky assets, is an acceptable position. To show closedness of the sets $\mathcal{A}_{t,k}^\rho(\omega)$, fix $\omega \in \Omega$ and consider a sequence $\{\zeta_n\} \subset \mathcal{A}_{t,k}^\rho(\omega)$ such that $\zeta_n \rightarrow \zeta$. Using Assumption 2.5 it holds that

$$\rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega) - K_t(\omega) = \lim_{n \rightarrow \infty} \rho_t(\mathfrak{E}(\zeta_n, t) - 1)(\omega) - K_t(\omega) \leq 0$$

holds for all $t \in [0, T]$ and which implies that $\zeta \in \mathcal{A}_t(\omega)$. The latter, together with the fact that $\zeta \in [-k, k]^n$ finalizes the proof. \square

Definition A.2. A correspondence ϕ between a measurable space (Θ, \mathcal{G}) and a topological space X is said to be *weakly measurable* if for all $F \subset X$ closed, the lower inverse of F , defined as

$$\phi^l(F) := \{\theta \in \Theta \mid \phi(\theta) \cap F \neq \emptyset\},$$

belongs to \mathcal{G} .

In the case of compact-valued correspondences, weak-measurability and Borel measurability (in terms of the Borel σ -algebra generated by the Hausdorff metric) are equivalent notions. Given a correspondence $\phi : \Omega \times [0, T] \mapsto \mathbb{R}^n$ we define the corresponding closure correspondence via $\bar{\phi}(\omega, t) := \overline{\phi(\omega, t)}$. For notational purposes let

$$f((t, \omega), \zeta) = \rho_t(\mathfrak{E}(\zeta, t) - 1)(\omega) - K_t(\omega).$$

Recall that \mathcal{P} denotes the predictable σ -algebra on $[0, T] \times \Omega$. The function $f((\cdot, \cdot), \cdot)$ is a Carathéodory function with respect to \mathcal{P} , i.e. it is continuous in ζ and \mathcal{P} -measurable in (t, ω) .

²For a comprehensive overview of the theory of measurable correspondences, we refer the reader to [1].

Proposition A.3. *For any $k \in \mathbb{N}$, the correspondence $\mathcal{A}_{\cdot,k}^\rho : [0, T] \times \Omega \rightarrow \mathbb{R}^n$ is weakly \mathcal{P} -measurable.*

Proof. Let $F \subset \mathbb{R}^n$ be closed and consider $\{\zeta_m\}_{m=1}^\infty \subset F$ dense. For $\eta \in \mathbb{N}$ let

$${}^\eta \mathcal{A}_{t,k}^\rho(\omega) := \left\{ \zeta \in [-k, k]^n \mid f((t, \omega), \zeta) < \frac{1}{\eta} \right\}.$$

We have that

$$\begin{aligned} ({}^\eta \mathcal{A}_{\cdot,k}^\rho)^l(F) &= \left\{ (t, \omega) \in [0, T] \times \Omega \mid f((t, \omega), \zeta) < \frac{1}{\eta} \text{ for some } \zeta \in F \right\} \\ &= \left\{ (t, \omega) \in [0, T] \times \Omega \mid f((t, \omega), \zeta_m) < \frac{1}{\eta} \text{ for some } m \in \mathbb{N} \right\} \\ &= \bigcup_{m=1}^\infty f^{-1}((\cdot, \cdot), \zeta_m) \left(-\infty, \frac{1}{\eta} \right). \end{aligned}$$

The second equality holds because f is continuous in ζ , $\{\zeta_m\}_{m=1}^\infty$ is dense and $(-\infty, 1/\eta)$ is open. Since f is Carathéodory, then $f^{-1}((\cdot, \cdot), \zeta) \left(-\infty, \frac{1}{\eta} \right) \in \mathcal{P}$, hence for all $\eta \in \mathbb{N}$, the correspondence ${}^\eta \mathcal{A}_{\cdot,k}^\rho$ is weakly \mathcal{P} -measurable. Next we have

$$\mathcal{A}_{t,k}^\rho(\omega) \subset \overline{{}^\eta \mathcal{A}_{t,k}^\rho(\omega)} \subset \left\{ \zeta \in [-k, k]^n \mid f((t, \omega), \zeta) \leq \frac{1}{\eta} \right\},$$

where the second inclusion follows again from the continuity of f in ζ . This implies that

$$\mathcal{A}_{t,k}^\rho(\omega) = \bigcap_{\eta=1}^\infty \overline{{}^\eta \mathcal{A}_{t,k}^\rho(\omega)}.$$

By Lemma 18.4 (part 3) in [1], we conclude that the correspondence $(t, \omega) \mapsto \mathcal{A}_{t,k}^\rho(\omega)$ is weakly \mathcal{P} -measurable. \square

The following theorem, whose proof can be found in [1], page 595, plays an important role in the proof of predictability of our BSDE's driver:

Theorem A.4. *A nonempty-valued correspondence mapping a measurable space into a separable, metrizable space is weakly-measurable if and only if its associated distance function is a Carathéodory function.*

Acknowledgment. The authors are very thankful to Jianing Zhang for his guidance concerning the numerical simulations.

References

1. Aliprantis, C. D. and Border, K. C.: *Infinite dimensional analysis*. Springer, Berlin, third edition, 2006. A hitchhiker's guide.
2. Atkinson, C. and Papakokkinou, M.: Theory of optimal consumption and portfolio selection under a capital-at-risk (CaR) and a value-at-risk (VaR) constraint, *IMA J. Manag. Math.* **16(1)** (2005) 37–70.
3. Basak, S. and Shapiro, A.: Value-at-risk-based risk management: optimal policies and asset prices, *Rev. Financial Studies* **14** (2001) 371–405.
4. Bender, C. and Denk, R.: A forward scheme for backward SDEs, *Stochastic Process. Appl.* **117(12)** (2007) 1793–1812.

5. Berkelaar, A. and Cumperayot, P. and Kouwenberg, R.: The effect of var-based risk management on asset prices and volatility smile, *European Financial Management* **16(1)** (2005) 65–78.
6. Cuoco, D. and He, H. and Isaenko, S.: Optimal dynamic trading strategies with risk limits, *Oper. Res.* **56(2)** (2008) 358–368.
7. Cvitanić, J. and Karatzas, I.: Convex duality in constrained portfolio optimization, *Ann. Appl. Probab.* **2(4)** (1992) 767–818.
8. Dmitrasinović-Vidović, G. and Lari-Lavassani, C., and Li, X., and Ware, X.: Dynamic portfolio selection under capital at risk *Journal of Probability and Statistics* (2010) 1–26.
9. dos Reis, G.: On some properties of solutions of quadratic growth BSDE and applications in finance and insurance, *PhD thesis, Humboldt University in Berlin*, <http://www.math.tu-berlin.de/~dosreis/publications/GdosReis-PhD-Thesis.pdf>, 2010.
10. Emmer, S. and Klüppelberg, C. and Korn, R.: Optimal portfolios with bounded capital at risk, *Math. Finance* **11(4)** (2001) 365–384.
11. Gabih, A., and Sass, J. and Wunderlich, R.: Utility maximization under bounded expected loss, *Stoch. Models* **25(3)** (2009) 375–407.
12. Gundel, A. and Weber, S.: Robust utility maximization with limited downside risk in incomplete markets, *Stochastic Process. Appl.* **117(11)** (2007) 1663–1688.
13. Hu, Y. and Imkeller, P. and Müller, M.: Utility maximization in incomplete markets, *Ann. Appl. Probab.* **15(3)** (2005) 1691–1712.
14. Imkeller, P. and Dos Reis, G.: Path regularity and explicit convergence rate for BSDE with truncated quadratic growth, *Stochastic Process. Appl.* **120(3)** (2010) 348–379.
15. Kazamaki, N.: *Continuous exponential martingales and BMO*, volume 1579 of *Lecture Notes in Mathematics*, Springer-Verlag, Berlin, 1994.
16. Klüppelberg, C. and Pergamenchtchikov, S.: Optimal consumption and investment with bounded downside risk for power utility functions. In *Optimality and risk—modern trends in mathematical finance*, pages 133–169. Springer, Berlin, 2009.
17. Kobylanski, M.: Backward stochastic differential equations and partial differential equations with quadratic growth, *Ann. Probab.* **28(2)** (2000) 558–602.
18. Leippold, M. and Trojani, F. and Vanini, P.: Equilibrium impact of value-at-risk regulation, *J. Econom. Dynam. Control* **30(8)** (2006) 1277–1313.
19. Morlais, M.-A.: Quadratic BSDEs driven by a continuous martingale and applications to the utility maximization problem, *Finance Stoch.* **13(1)** (2009) 121–150.
20. Nutz, M.: The opportunity process for optimal consumption and investment with power utility, *Math. Financ. Econ.* **3(3-4)** (2010) 139–159.
21. Pirvu, T. A.: Portfolio optimization under the value-at-risk constraint, *Quant. Finance* **7(2)** (2007) 125–136.
22. Pirvu, T. A. and Žitković, G.: Maximizing the growth rate under risk constraints, *Math. Finance* **19(3)** (2009) 423–455.
23. Prieto, R.: Dynamic Equilibrium with Heterogeneous Agents and Risk Constraints, *Preprint* (2010).
24. Putschögl, W.G. and Sass, J.: Optimal investment under dynamic risk constraints and partial information, *Quant. Finance* **11(10)** (2011) 1547–1564.
25. Tsukahara, H.: One-parameter families of distortion risk measures, *Math. Finance* **19(4)** (2009) 691–705.
26. Yiu, K. F. C.: Optimal portfolios under a value-at-risk constraint, *J. Econom. Dynam. Control* **28(7)** (2004) 1317–1334.

SANTIAGO MORENO-BROMBERG: INSTITUT FÜR BANKING UND FINANCE, UNIVERSITÄT ZÜRICH,
PLATTENSTR. 32, 8032 ZÜRICH, SWITZERLAND

E-mail address: santiago.moreno@bf.uzh.ch

URL: http://www.bf.uzh.ch/cms/institut/_165-1635.html

TRAIAN A. PIRVU: MATHEMATICS AND STATISTICS DEPARTMENT, MCMMASTER UNIVERSITY, HAMIL-
TON, ON, CANADA

E-mail address: tpirvu@math.mcmaster.ca

URL: <http://www.math.mcmaster.ca/~tpirvu/>

ANTHONY RÉVEILLAC: CEREMADE UMR CNRS 7534, UNIVERSITÉ PARIS DAUPHINE, PLACE DU
MARÉCHAL DE LATTRE DE TASSIGNY, 75775 PARIS CEDEX 16, FRANCE

E-mail address: anthony.reveillac@ceremade.dauphine.fr

URL: <http://www.ceremade.dauphine.fr/~areveill/index.htm>