M4056 Lecture Notes.

Recall:

- If X is $n(0,1), M_X(t) = e^{t^2/2}$
- $M_{aX+b}(t) = e^{bt} M_X(at)$

If Y is $n(\mu, \sigma^2)$, then $Y = \sigma X + \mu$, where X is n(0, 1). Thus,

if
$$Y \sim n(\mu, \sigma^2)$$
, $M_Y(t) = e^{\mu t + \sigma^2 t^2/2}$.

Suppose X_1, \ldots, X_n are any random variables. Then

$$M_{\overline{X}}(t) = E e^{t\overline{X}} = E e^{(t/n)(X_1 + \dots + X_n)} = M_{X_1 + \dots + X_n}(t/n).$$

We have seen previously that the mgf of distribution of a sum of independent random variables is the product of the mfgs of the summands. Hence:

if
$$X_1, \ldots, X_n$$
 are *iid* $\sim X$, then $M_{\overline{X}}(t) = [M_X(t/n)]^n$.

(This appears in your book as **Theorem 5.2.7**.)

This enables us to calculate the distribution of the sum of *iid* normal variables. If X_1, \ldots, X_n are *iid* $n(\mu, \sigma^2)$, then

$$M_{\overline{X}}(t) = \left[e^{\mu (t/n) + \sigma^2 (t/n)^2/2}\right]^n$$
$$= e^{n[\mu t/n + \sigma^2 (t/n)^2/2]}$$
$$= e^{\mu t + (\sigma^2/n) (t^2/2)}$$

The last expression is the mgf of an $n(\mu, \sigma^2/n)$ random variable. This is consistent with the observations we made last time: the expected value and variance of \overline{X} are as we have already seen they must be. But even more: the sample mean is itself normal, so these two numbers tell the whole story. We restate this important result:

Fact. If X_1, \ldots, X_n are *iid* $n(\mu, \sigma^2)$, then \overline{X} is $n(\mu, \sigma^2/n)$. (5.3.1.b)

Next, we turn to the distribution of the sample variance S^2 when samples of size n are drawn from an $n(\mu, \sigma^2)$ distribution.

Fact. \overline{X} and S^2 are independent. (5.3.1.a)

Fact. $(n-1)S^2/\sigma^2$ has a χ^2_{n-1} (chi squared with n-1 degrees of freedom) distribution. (5.3.1.c)